

# Repeated Measures Design

A decorative graphic consisting of several horizontal lines of varying colors (dark blue, light blue, and white) that extend across the width of the slide, positioned below the title.

# Repeated Measures ANOVA

- Instead of having one score per subject, experiments are frequently conducted in which multiple scores are gathered for each case
- Repeated Measures or Within-subjects design

# When to Use

- Measuring performance on the same variable over time
  - for example looking at changes in performance during training or before and after a specific treatment
- The same subject is measured multiple times under different conditions
  - for example performance when taking Drug A and performance when taking Drug B
- The same subjects provide measures/ratings on different characteristics
  - for example the desirability of different characteristics of interpersonal relationships
- Note how we could do some RM as regular between subjects designs
  - Ex. Randomly assign to drug A or B

# Advantages to RM design

- Design – nonsystematic variance (i.e. error, that not under experimental control) is reduced
  - Take out error variance that's due to individual differences, resulting in more sensitivity/power for the treatment main effect
- Efficiency – fewer subjects are required

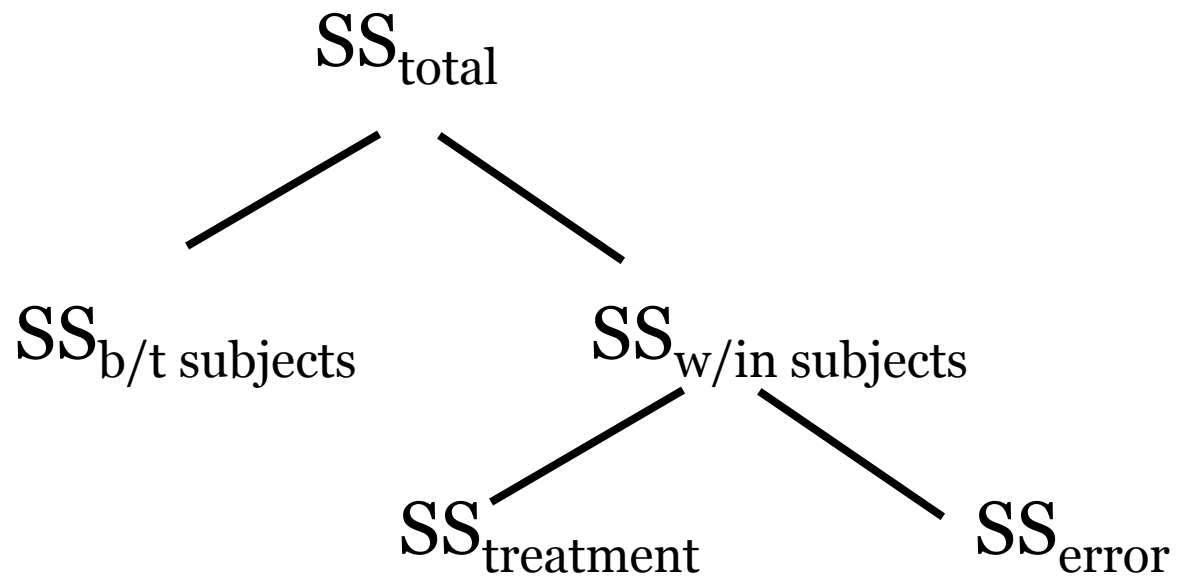
# Independence in ANOVA

- Analysis of variance as discussed previously assumes cells are independent
- But here that is not going to be the case
  - For example, those subjects who perform best in one condition are likely to perform best in the other conditions
- One thing to note though, observations still need to be independent from one case to the next

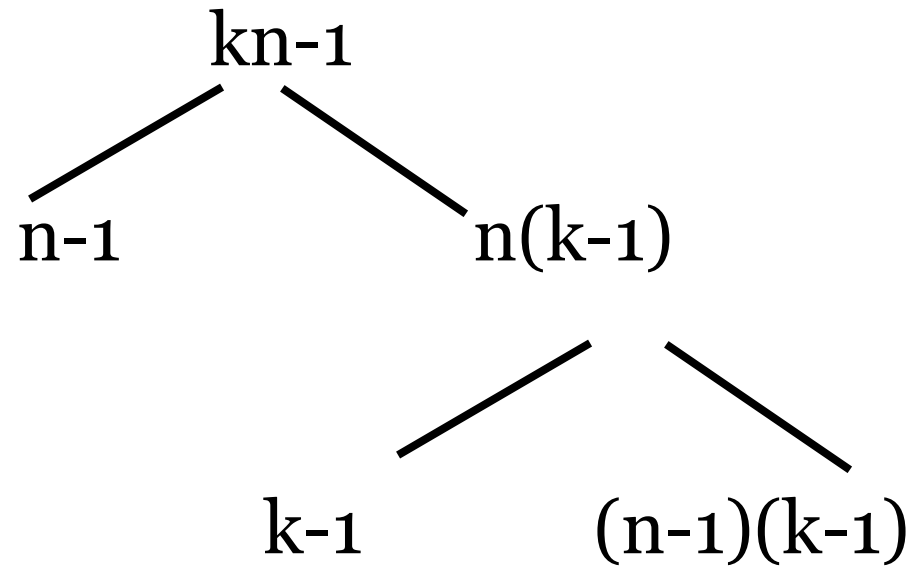
# Partialling out dependence

- Our differences of interest now reside within subjects and we are going to partial out differences between the subjects
  - This removes the dependence on subjects that causes the problem mentioned
- For example:
  - Subject 1 scores 10 in condition A and 14 in condition B
  - Subject 2 scores 6 in condition A and 10 in condition B
- In essence, what we want to consider is that both subjects score 2 less than their own overall mean score in condition A and 2 more than their own overall mean score in condition B

# Partition of SS



# Partitioning the degrees of freedom



# Understanding the Uncertainty

- $SS_{\text{total}}$ 
  - Deviation of each individual score from the grand mean
- $SS_{\text{b/t subjects}}$ 
  - Deviation of subjects' individual means (across treatments) from the grand mean.
  - In the RM setting, this is typically deemed uninteresting, as we can pretty much assume that 'subjects differ'<sup>1</sup>
- $SS_{\text{w/in subjects}}$ : How Ss vary about their own mean, breaks down into:
  - $SS_{\text{treatment}}$ 
    - As in between subjects ANOVA, is the comparison of treatment/measure means to each other (by examining their deviations from the grand mean)
    - However this is now a partition of the within subjects variation
  - $SS_{\text{error}}$ 
    - Variability of individuals' scores about their treatment/measure mean

# Example

- Effectiveness of mind control for different drugs

| Prisoner | Placebo  | DrugA    | DrugB    | DrugC    | Mean     |
|----------|----------|----------|----------|----------|----------|
| Henchman | 3        | 4        | 6        | 7        | <b>5</b> |
| Cobb     | 0        | 3        | 3        | 6        | <b>3</b> |
| Dutton   | 2        | 1        | 4        | 5        | <b>3</b> |
| The Rook | 0        | 1        | 3        | 4        | <b>2</b> |
| #6       | 0        | 1        | 4        | 3        | <b>2</b> |
| Mean     | <b>1</b> | <b>2</b> | <b>4</b> | <b>5</b> | <b>3</b> |

# Calculating SS

- Calculate  $SS_{\text{within}}$
- Conceptually it reflects the subjects' scores variation about their own individual means
- $(3-5)^2 + (4-5)^2 \dots (4-2)^2 + (3-2)^2 = 58$
- This is what will be broken down into treatment and error variance

# Calculating SS

- Calculate  $SS_{\text{treat}}$
- Conceptually it is the sum of the variability due to all treatment *pairs*
  - If we had only two treatments, the F for this would equal  $t^2$  for a paired samples t-test

- $SS_{\text{b/t}} = SS_{\text{treat}} = \frac{n \sum (\bar{T} - \bar{T}')^2}{k}$

- $SS_{\text{treat}} = n \sum (\bar{T} - \bar{G})^2$

- $SS_{\text{treat}} = 5[(1-3)^2 + (2-3)^2 + (4-3)^2 + (5-3)^2] = 50$

↑  
5 people in each treatment

↑ ↑ ↑ ↑  
Treatment means

↑  
Grand mean

# Calculating SS

- $SS_{\text{error}}$
- Residual variability
  - Unexplained variance, which includes subject by treatment interaction\*
- Recall that  $SS_{\text{w/in}} = SS_{\text{treat}} + SS_{\text{error}}$
- $SS_{\text{error}} = SS_{\text{w/in}} - SS_{\text{treat}}$
- $58 - 50 = 8$

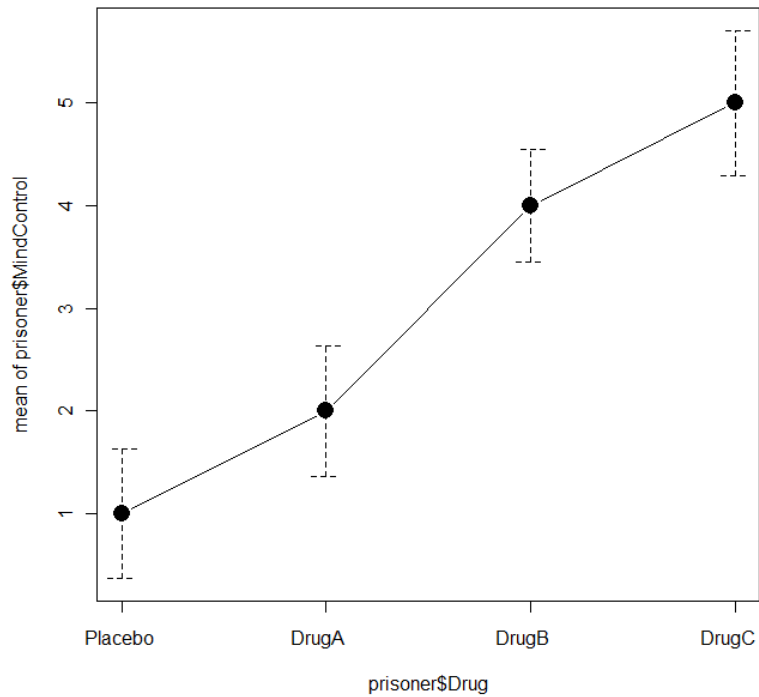
Tests of Within-Subjects Effects

Measure: MEASURE\_1

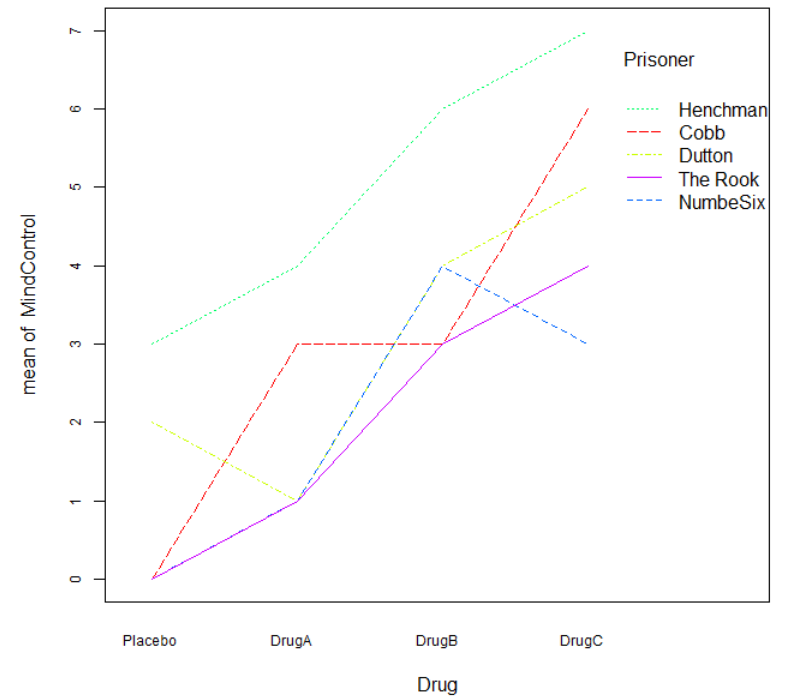
| Source      | Type III Sum of Squares | df | Mean Square | F      | Sig. | Partial Eta Squared |
|-------------|-------------------------|----|-------------|--------|------|---------------------|
| drug        | 50.000                  | 3  | 16.667      | 25.000 | .000 | .862                |
| Error(drug) | 8.000                   | 12 | .667        |        |      |                     |

# Plot of Effects

Plot of Means



Mind Control Effectiveness



# Effect size

Tests of Within-Subjects Effects

Measure: MEASURE\_1

| Source      | Type III Sum of Squares | df | Mean Square | F      | Sig. | Partial Eta Squared |
|-------------|-------------------------|----|-------------|--------|------|---------------------|
| drug        | 50.000                  | 3  | 16.667      | 25.000 | .000 | .862                |
| Error(drug) | 8.000                   | 12 | .667        |        |      |                     |

- Note that as we have partialled out error due to subject differences from the overall residual error, our measure of effect here is
  - $SS_{\text{effect}} / (SS_{\text{effect}} + SS_{\text{error}})$
- So this is PES not simply  $\eta^2$  as it was for one-way between subjects ANOVA
- $PES = 50/58$

# Interpretation

- As with a regular one-way Anova, the omnibus RM analysis tells us that there is some difference among the treatments (drugs)
- Often this is not a very interesting outcome, or at least, not where we want to stop in our analysis
- In this example we might want to know which drugs are better than which

# Contrasts and Multiple Comparisons

- If you had some particular relationship in mind you want to test due to theoretical reasons (e.g. a linear trend over time) one could test that by doing contrast analyses.
- This table compares standard contrasts available in statistical packages
  - Deviation, simple, difference, Helmert, repeated, and polynomial
- Given the nature of RM design, it is unlikely you wouldn't at least test for a trend of some kind

|                              |  |
|------------------------------|--|
| Deviation                    | Compares the mean of one level to the mean of all levels (grand mean)                              |
| Simple                       | Compares each mean to some reference mean (either the first or last category e.g. a control group) |
| Difference (reverse Helmert) | Compares level 1 to 2, level 3 with the mean of the previous two etc.                              |
| Helmert                      | Compares level 1 with all later, level 2 with the mean of all later, level 3 etc.                  |
| Repeated                     | Compares level 1 to level 2, level 2 to level 3, 3 to 4 and so on                                  |
| Polynomial                   | Tests for trends (e.g. linear) across levels   |

# Multiple comparisons

- However with our drug example we are not dealing with a time based model and may not have any preconceived notions of what to expect<sup>1</sup>
- So now how are you going to do conduct a post hoc analysis?
- Technically you could flip your data so that treatments are in the rows with their corresponding score, run a regular one-way ANOVA, and do Tukey's etc. as part of your analysis.
- However you would still have problems because the appropriate error term would not be used in the analysis.
  - B/t subjects effects not removed from error term

# Multiple comparisons

- The process for doing basic comparisons remains the same
- In the case of repeated measures, as Howell notes (citing Maxwell) one may elect in this case to test them separately as opposed to using a pooled error term<sup>1</sup>
  - The reason for doing so is that such tests would be extremely sensitive to departures from the sphericity assumption
- However using the  $MS_{\text{error}}$  we can test multiple comparisons and once you have your resulting t-statistic, one can get the probability associated with that

$$\sqrt{\frac{\sum(a^2)MS_{\text{error}}}{n}}$$

# Multiple comparisons

- Example in R comparing placebo and Drug A
  - $t = 1.93$
  - $\text{pt}(1.93, 4, \text{lower.tail}=\text{F})$
  - .063 one-tailed or .126 two-tailed

$$\frac{\psi}{\sqrt{\frac{\sum(a^2)MS_{error}}{n}}}$$
$$\frac{1}{\sqrt{\frac{2 * .667}{5}}} = 1.93$$

$MS_{error}$  from ANOVA table

# Multiple comparisons

- While one could correct in Bonferroni fashion, there is a False discovery rate for dependent tests
  - Will control for overall type I error rate among the rejected tests
- The following example uses just the output from standard pairwise ts for simplicity

# Multiple comparisons

- Output from t-tests

|        |                 | Paired Differences |                |                    |   |        | t      | df | Sig. (2-tailed) |
|--------|-----------------|--------------------|----------------|--------------------|---|--------|--------|----|-----------------|
|        |                 | Mean               | Std. Deviation | Std. Error<br>Mean | 95% Confidence<br>Interval of the<br>Difference |        |        |    |                 |
|        |                 |                    |                |                    | Lower   | Upper  |        |    |                 |
| Pair 1 | Placebo - DrugA | -1.000             | 1.414          | .632               | -2.756  | .756   | -1.581 | 4  | .189            |
| Pair 2 | Placebo - DrugB | -3.000             | .707           | .316               | -3.878  | -2.122 | -9.487 | 4  | .001            |
| Pair 3 | Placebo - DrugC | -4.000             | 1.225          | .548               | -5.521  | -2.479 | -7.303 | 4  | .002            |
| Pair 4 | DrugA - DrugB   | -2.000             | 1.225          | .548               | -3.521  | -.479  | -3.651 | 4  | .022            |
| Pair 5 | DrugA - DrugC   | -3.000             | .707           | .316               | -3.878  | -2.122 | -9.487 | 4  | .001            |
| Pair 6 | DrugB - DrugC   | -1.000             | 1.414          | .632               | -2.756  | .756   | -1.581 | 4  | .189            |

# Multiple comparisons

- Output from R. P-values from previous slide have been ordered

|      | rawp  | Bonferroni | Holm  | Hochberg | SidakSS    | SidakSD     | BH    | BY      |
|------|-------|------------|-------|----------|------------|-------------|-------|---------|
| [1,] | 0.001 | 0.006      | 0.006 | 0.005    | 0.00598502 | 0.005985020 | 0.003 | 0.00735 |
| [2,] | 0.001 | 0.006      | 0.006 | 0.005    | 0.00598502 | 0.005985020 | 0.003 | 0.00735 |
| [3,] | 0.002 | 0.012      | 0.008 | 0.008    | 0.01194016 | 0.007976032 | 0.004 | 0.00980 |
| [4,] | 0.022 | 0.132      | 0.066 | 0.066    | 0.12494948 | 0.064558648 | 0.033 | 0.08085 |
| [5,] | 0.189 | 1.000      | 0.378 | 0.189    | 0.71547193 | 0.342279000 | 0.189 | 0.46305 |
| [6,] | 0.189 | 1.000      | 0.378 | 0.189    | 0.71547193 | 0.342279000 | 0.189 | 0.46305 |

- The last column is the Benjamini and Yekutieli correction of the p-value that takes into account the dependent nature of our contrasts
- A general explanation might lump Drug A as ineffective (not statistically different from the placebo), and B & C similarly effective

# Assumptions for RM

- Standard ANOVA assumptions
  - Homogeneity of variances
  - Normality
  - Independent observations
- For RM design we are looking for homogeneity of covariances among the treatments e.g.  $t_1, t_2, t_3$ 
  - Special case of HoV
  - Sphericity
    - When the variance of the difference scores for any pair of groups is the same as for any other pair

# Sphericity

- Observations may covary across time, dose etc., and we would expect them to. But the degree of covariance must be similar.
  - If covariances are heterogeneous, the error term will generally be an underestimate and F tests will be positively biased
- Such circumstances may arise due to carry-over effects, practice effects, fatigue and sensitization

# Sphericity

- Suppose the repeated measure factor of TIME had 3 levels – before, after and follow-up scores for each individual
- RM ANOVA assumes that the 3 correlations
  - $r$  ( Before-After )
  - $r$  ( Before-Follow up )
  - $r$  ( After-Follow up )
- Are all about the same in size
  - i.e. any difference due to sampling error

# Sphericity

- Tests can be run to show if they are not
  - Mauchly's test of Sphericity generates a significant chi square
  - Again, when testing assumptions we typically hope they do not return a significant result
- A correction factor called EPSILON is applied to the degrees of freedom of the error term when calculating the significance of F
  - This is default output for some statistical programs

# Sphericity

- If the Mauchly Sphericity test is significant, then use the Corrected significance value for F
  - Different approaches will apply different corrective factors<sup>1</sup>
- If there are only 2 levels of the factor, why is sphericity is not a problem?

# More Examples



# Another RM example

| Before | During | After |
|--------|--------|-------|
| 14     | 15     | 14    |
| 34     | 33     | 22    |
| 26     | 28     | 25    |
| 30     | 33     | 18    |
| 35     | 33     | 30    |
| 16     | 31     | 23    |
| 27     | 28     | 28    |
| 28     | 33     | 19    |
| 28     | 34     | 27    |
| 25     | 27     | 25    |

- Students were asked to rate their stress on a 50 point scale in the week before, the week of, or the week after their midterm exam

# Analysis

- For comparison, first analyze these data as if they were from a between subjects design
- Then conduct the analysis again as if they come from a repeated measures design

# Comparing outputs

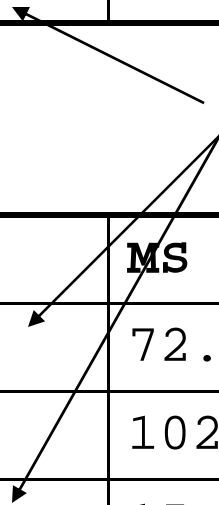
- Between-subjects

| Source | df | SS    | MS     | F     | p     |
|--------|----|-------|--------|-------|-------|
| time   | 2  | 204.8 | 102.4  | 2.981 | .0676 |
| error  | 27 | 927.5 | 34.352 |       |       |

- Within subjects

| Source  | df | SS    | MS     | F     | p     |
|---------|----|-------|--------|-------|-------|
| Subject | 9  | 654.3 | 72.7   |       |       |
| time    | 2  | 204.8 | 102.4  | 6.747 | .0065 |
| error   | 18 | 273.2 | 15.178 |       |       |

+



# Comparing outputs

- SS due to subjects has been discarded from the error term in the analysis of the treatment in the RM design
- Same b/t treatment effect
- Less in error term
- More power in analyzing as repeated measures
- Given outliers, robust check is below using trimmed means

```
[1] "The number of groups to be compared is"
```

```
[1] 3
```

```
$test
```

```
[1] 7.11542
```

```
$df
```

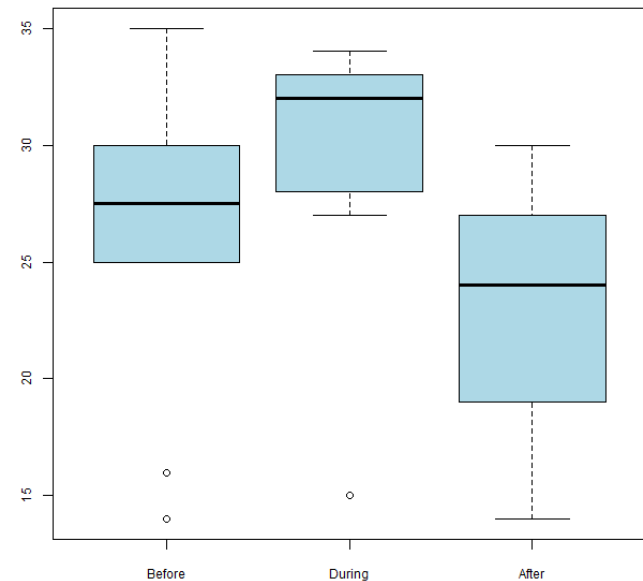
```
[1] 1.229382 6.146911
```

```
$siglevel
```

```
[1] 0.03272527
```

# Trend Contrast

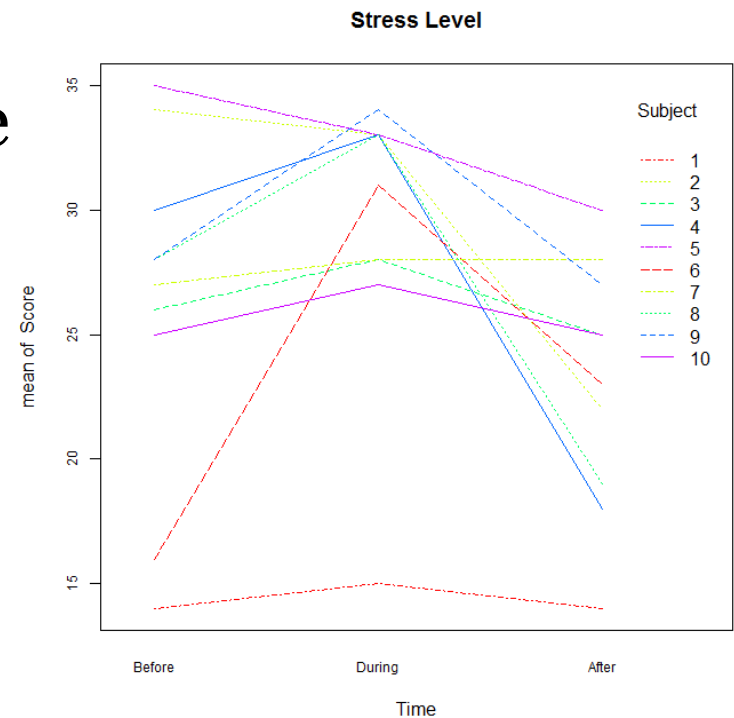
- It would be unlikely to no expect differences in a certain pattern, in particular, we'd expect a hump like the one portrayed in the data
- Testing for a quadratic trend would tell us if in fact this were the case



# Trend Contrast

- Plot of subject differences
- A quadratic trend does appear best

|                 | Df | Sum Sq  | Mean Sq | F value | Pr(>F)      |
|-----------------|----|---------|---------|---------|-------------|
| Time            | 2  | 204.800 | 102.400 | 6.7467  | 0.006508 ** |
| Time: Linear    | 1  | 51.200  | 51.200  | 3.3734  | 0.082830 .  |
| Time: Quadratic | 1  | 153.600 | 153.600 | 10.1201 | 0.005172 ** |
| Residuals       | 18 | 273.200 | 15.178  |         |             |



## Another example (from Howell)

- An experiment designed to look at the effects of a relaxation technique on migraine headaches
- 9 subjects are asked to record the number of hours of such headaches each week to provide a baseline, then are given several weeks of training in relaxation technique – during which the number of hours of migraines per week were also recorded
- A single within-subject effect with 5 levels (week 1, week 2, week 3, week 4, week 5)

# Data

| baseline |       | training |       |       |
|----------|-------|----------|-------|-------|
| week1    | week2 | week3    | week4 | week5 |
| 21       | 22    | 8        | 6     | 6     |
| 20       | 19    | 10       | 4     | 4     |
| 17       | 15    | 5        | 4     | 5     |
| 25       | 30    | 13       | 12    | 17    |
| 30       | 27    | 13       | 8     | 6     |
| 19       | 27    | 8        | 7     | 4     |
| 26       | 16    | 5        | 2     | 5     |
| 17       | 18    | 8        | 1     | 5     |
| 26       | 24    | 14       | 8     | 9     |

# Results

| Source  | df | Sum of Squares | Mean Square | F      | p     |
|---------|----|----------------|-------------|--------|-------|
| Subject | 8  | 486.711        | 60.839      |        |       |
| time    | 4  | 2449.200       | 612.300     | 85.042 | .0001 |
| error   | 32 | 230.400        | 7.200       |        |       |

Note however that we really had 2 within subjects variables condition (baseline vs. training) and time (week 1-5)

How would we analyze that?

Tune in next week! Same bat time! Same bat channel!

# Appendix

More stuff that likely won't make  
the cut at next time

# Review - Covariance

- Measure the joint variance of two (or more) variables
- Use cross product of deviation of scores from their group mean
- Formula:

$$\text{covar}_{xy} = \frac{\sum ((X - \bar{X})(Y - \bar{Y}))}{n - 1}$$

$$r = \frac{\text{covar}_{xy}}{s_x s_y}$$

# Correcting for deviations

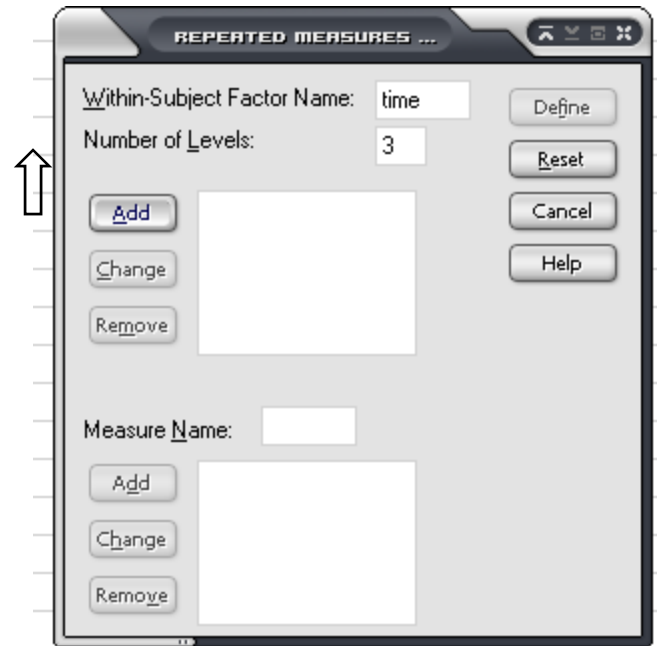
- Epsilon ( $\varepsilon$ ) measures the degree to which covariance matrix deviates from compound symmetry
  - All the variances of the treatments are equal and covariances of treatments are equal
- When  $\varepsilon = 1$  then matrix is symmetrical, and normal df apply
- When  $\varepsilon = k^{-1}$  then matrix has maximum heterogeneity of variance and the F ratio should have 1, n-1 df

# Estimating F

- Two different approaches, adjusting the df
  - Note that the F statistic will be the same, but the df will vary
- Conservative: Box's/Greenhouse-Geisser
- Liberal: Huynh-Feldt
- Huynh-Feldt tends to overestimate  $\varepsilon$  (can be  $> 1$ , at which point it is set to 1)
- Some debate, but Huynh-Feldt is probably regarded as the method of choice, see Glass & Hopkins

# Conducting RM in SPSS

- “Analyze” ⇒ “General Linear Model” ⇒ “Repeated Measures”
- Type “time” into the within-subject factor name text box (in place of “factor1”)
- Type “3” in the Number of Levels box
- Click “Add”
- Click “Define”



# Conducting RM in SPSS

- Move the variables related to the treatment to the within subjects box
- Select other options as desired

